

## Bayesian Inference

Recommended text:  
D.S. Sivia & J. Skilling  
"Data Analysis: A Bayesian Tutorial"

Problem: Have some data  $\underline{s}[t] = \{s^0, s^1, \dots\}$

- Data is noisy.

- Assume  $\underline{s}[t] = \underline{n}[t] + \underline{h}[t]$

•  $\underline{n}[t]$ : noise

•  $\underline{h}[t] = \{h(t_j; \lambda)\}$  signal, given by model  $h(t_j; \lambda)$ .

The signal depends on a set of parameters  $\lambda$ .

Ex: a CW from a binary,  $\lambda = \{m_1, m_2, D_L, \dots\}$

We want to know:

What is the probability the signal has params  $\lambda'$ , given our data & signal model;

$$p(\lambda' | \underline{s}, h) = ?$$

↑      ↴  
params.    given

Bayes Theorem:

$$p(\underline{\lambda}' | \underline{s}, h) = \frac{p(\underline{s} | \underline{\lambda}', h) p(\underline{\lambda}' | h)}{p(\underline{s} | h)}$$

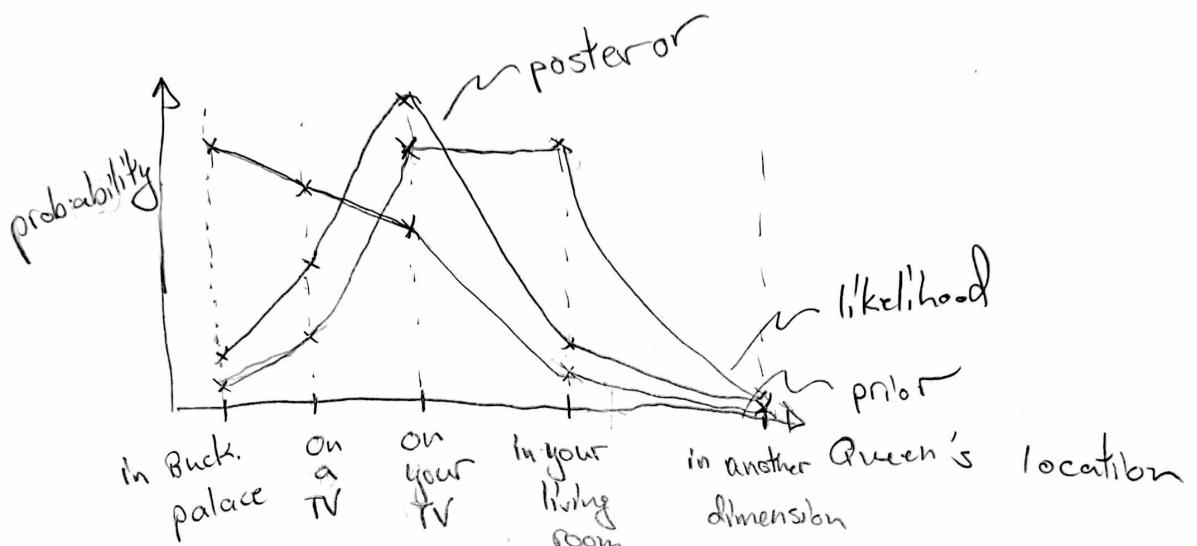
"likelihood"      "prior"  
 ↓                    ↓  
 P                    R  
 "posterior"        "evidence"

- likelihood: prob. of observing the data assuming the signal has params  $\underline{\lambda}'$
- prior: our prior belief that a signal can have parameters  $\underline{\lambda}'$
- evidence: normalization constant, used for comparing models.



Where is the Queen of England?

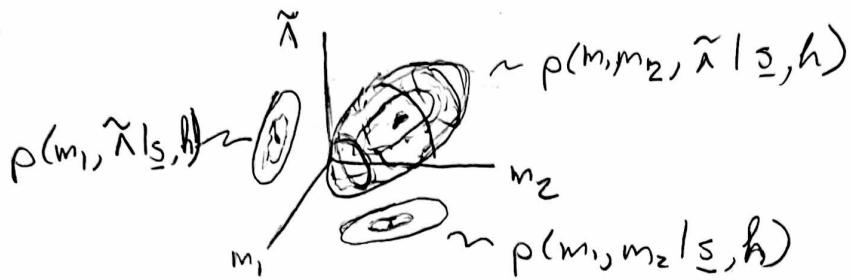
Given: you hear the Queen's voice in your living room



## Marginalization

Full posterior prob. is multidimensional. Ex: for BBH,

$\underline{\lambda}$  has 15 params  $\Rightarrow$  15 dimensional param. space.



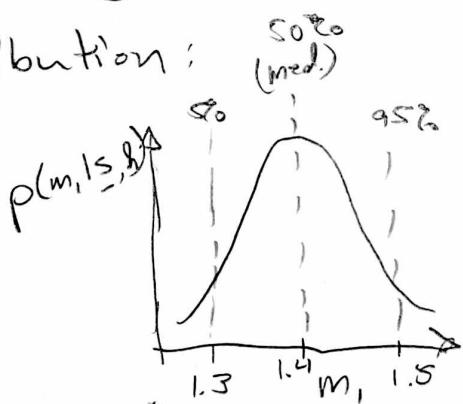
Integrating posterior over a subset of params yields marginal:

$$p(\{x^0, \dots, x^k\} | s, h) = \int p(\underline{x} | s, h) d\underline{x}^{k+1} \dots d\underline{x}^n$$

$$= \frac{1}{p(s|h)} \int p(s|\underline{x}, h) p(\underline{x} | h) d\underline{x}^{k+1} \dots d\underline{x}^n$$

Integrating over all but 1 param yields a 1D

distribution:



Quote credible intervals  
using this

$$\Delta m_1 = 1.4^{+0.1}_{-0.1} M_\odot$$

## Evidence

Since post. is a pdf, marginalizing over all params:

$$\int p(\lambda | s, h) d\lambda = 1$$

$$\Rightarrow \boxed{p(s | h) = \int p(s | \lambda, h) p(\lambda | h) d\lambda}$$

Evidence is used for model selection. Say we have 2 different signal models  $h_A, h_B$ .

## Bayes Factor

$$\boxed{\beta(A, B | s) = \frac{p(s | h_A)}{p(s | h_B)}}$$

$\log \beta > 0 \rightarrow$  model A favored

$\log \beta < 0 \rightarrow$  model B favored

$\log \beta = 0 \rightarrow$  can't tell

Note: if only interested in posterior for a single model, sufficient to evaluate

$$p(\lambda | s, h) \propto p(s | \lambda, h) p(\lambda | h)$$

In order to evaluate posterior, we need  
a signal model  $h$  and a noise model  $n$ .

- $h$  comes from GR (... & other physics)
- What is the noise model?

## Derivation of Likelihood

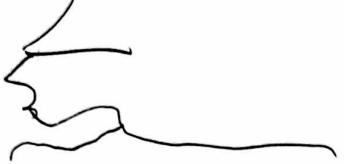
L.S. Finn, PRD 46, S236-S249  
(1992)

Wainstein & Zubakov, Extraction  
of signals from noise, 1962

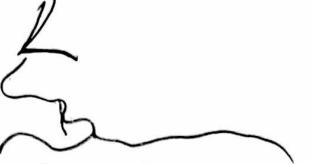
Have network of  $K$  AW detectors ( $K=3$ , HLV).

Sample their output at rate  $1/\Delta t$  for time  $T$ ,  
get  $N = \lfloor T/\Delta t \rfloor$  samples

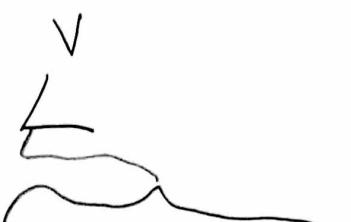
H



L



V



$$\underline{s}_{\text{Net}} = \{s_H^0, s_H^1, \dots, s_H^N, s_L^0, s_L^1, \dots, s_L^N, s_V^0, s_V^1, \dots, s_V^N\}$$

Consider the hypothesis that the data contain no signal,  $\underline{s} = \underline{n}$ . What is the prob. of observing the given realization?  $p(\underline{s}_{\text{Net}} | \underline{n}) = ?$

Assume noise is a stochastic Gaussian process

$$p(\underline{s}_{\text{Net}} | \underline{n}) = \frac{\exp\left[-\frac{1}{2} \underline{s}_{\text{Net}}^T \underline{\Sigma}_{\text{Net}}^{-1} \underline{s}_{\text{Net}}\right]}{(2\pi)^{NK} \det \underline{\Sigma}_{\text{Net}}}^{\frac{1}{2}}$$

$\underline{\Sigma}_{\text{Net}}$  is the covariance matrix of the noise

Assuming 0 mean, is:

$$\underline{\Sigma}_{\text{Net}}[j,k] = \langle \underline{s}[j] \underline{s}[k] \rangle, \quad \langle \dots \rangle \text{ is an ensemble average}$$

$$\underline{\Sigma}_{\text{net}} = \begin{pmatrix} \text{HH} & \text{HL} & \text{HV} \\ \text{LH} & \text{LL} & \text{LV} \\ \text{VL} & \text{VL} & \text{VW} \end{pmatrix}$$

Assume noise is uncorrelated between detectors

→  $\underline{\Sigma}_{\text{net}}$  block diagonal

$$p(s_{\text{net}} | n) = \frac{\exp\left[-\frac{1}{2} \sum_{d=1}^K \underline{s}_d^T \underline{\Sigma}_d^{-1} \underline{s}_d\right]}{(2\pi)^{NK} \prod_{d=1}^K \det \Sigma_d^{1/2}}$$

→ Need  $\underline{\Sigma}_d^{-1}$  (will drop detector subscripts from here)

=

$$\Sigma[j, k] = \langle \underline{s}[j] \underline{s}[k] \rangle$$

$$= \langle \underline{s}[j] \underline{s}[\Delta_{kj} + j] \rangle, \quad \Delta_{kj} = k - j$$

$$= \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{l=0}^{n-1} \underline{s}^{(l)}[j] \underline{s}^{(l)}[\Delta_{kj} + j]$$

→ In general cov. depends on both time considered  $t_j = j \Delta t$   
& displacement  $\mathcal{I}_{kj} = \Delta_{kj} \Delta t$

Assume noise is Wide sense stationary (WSS)

WSS: both mean & variance is constant

Then we can add any const. to both indices & get the same result:

$$\begin{aligned}\sum [e_{j+m}, k+m] &= \sum [e_j, k] \\ &= \langle s[0], s[\Delta_{kj}] \rangle \\ &= \langle s[-\Delta_{kj}], s[0] \rangle\end{aligned}$$

$\Rightarrow \Sigma$  is  $\pm$ -symmetric

- diagonals are equal ("Toeplitz")
- elements are even functions of  $\Delta_{kj}$

$$\Sigma = \begin{pmatrix} \cdot & & & \\ \cdot & \cdot & \cdot & \\ \cdot & \cdot & \cdot & \\ \cdot & \cdot & \cdot & \cdot \end{pmatrix}$$

=

Elements are ensemble averages.

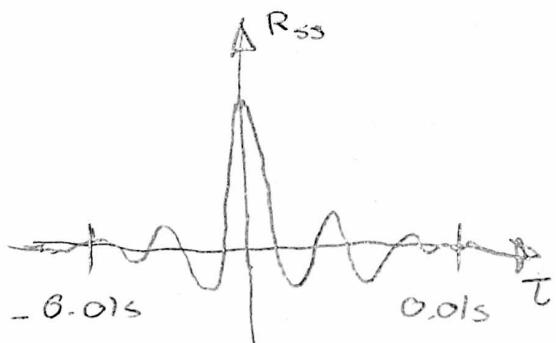
Assume data is ergodic  $\rightarrow$  get new realizations via time

Can replace ensemble average with time average:

$$\begin{aligned}
 \sum_{j,k} [f_j, k] &= \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{l=0}^{n-1} S^{(l)}[0] S^{(l)}[\Delta_{kj}] \\
 &= \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{l=0}^{n-1} S[l] S[\Delta_{kj} + l] \\
 &= \lim_{n \rightarrow \infty} \frac{1}{2n} \sum_{l=-n}^{n-1} S[l] S[\Delta_{kj} + l] \quad \text{even func.} \\
 &= \frac{1}{2} R_{ss}((k-j) \Delta t)
 \end{aligned}$$

$R_{ss}(\tau)$  is the autocorrelation function of the data.

For AW detectors  $R_{ss}$  looks like:



If  $R_{ss}(\tau) \rightarrow 0$  in some finite time  $T_{max}$  then all diagonals with  $|\Delta_{kj}| > [T_{max}/\Delta t] = \Delta_{max} = 0$

In this case, cov. matrix looks like:

$$\underline{\Sigma} = \frac{1}{2} \begin{pmatrix} R_{ss}[0] & R_{ss}[\Delta_{\max}] & & \\ R_{ss}[\Delta_{\max}] & 0 & & \\ & & R_{ss}[\Delta_{\max}] & \\ & & & R_{ss}[0] \end{pmatrix}$$

This is similar to:

$$\underline{C} = \frac{1}{2} \begin{pmatrix} R_{ss}[0] & R_{ss}[\Delta_{\max}] & R_{ss}[1] & & \\ R_{ss}[\Delta_{\max}] & 0 & & & \\ & & R_{ss}[\Delta_{\max}] & & \\ & & & R_{ss}[1] & \\ & & & & R_{ss}[0] \end{pmatrix}$$

$\underline{C}$  is a circulant matrix; i.e., has form

$$\begin{pmatrix} a & b & c & d \\ b & a & b & c \\ c & b & a & b \\ d & c & b & a \end{pmatrix} \Rightarrow \underline{C}[j+N] = \underline{C}[j]$$

=

\* All circulant matrices have the same eigenvectors:

$$u_p[k] = \frac{1}{\sqrt{N}} e^{-2\pi i k p / N}$$

See R.M. Gray,  
 "Toeplitz & Circulant  
 Matrices: A Review"

12

Not true of Toeplitz matrices (like  $\underline{\Sigma}$ ), but b/c  $\underline{\Sigma}$  is a banded Toeplitz, it is asymptotically equivalent to  $C$ , i.e.:

$$\lim_{N \rightarrow \infty} |\underline{\Sigma} - C| = 0$$

This means: we can find eigenvalues of  $\underline{\Sigma}$  using

$$\underline{\Sigma}_{UP} \approx \lambda_p \underline{u}_p$$

as long as  $\Delta_{\max} \ll [N/2]$ .

Solving, get:

$$\lambda_p = \frac{1}{2} \operatorname{Re} \left\{ \sum_{l=-N/2}^{N/2} R_{ss}[l] e^{-2\pi i pl/N} \right\}$$

)  $R_{ss}$  even

$$= \frac{1}{2} \operatorname{Re} \left\{ \sum_{l=0}^{N-1} \dots \right\}$$

$$= \frac{1}{2} \operatorname{Re} \left\{ \tilde{R}_{ss}[p] / \delta t \right\}$$

The  $\operatorname{Re}$  arises  
b/c  $\underline{\Sigma}$  is sym.  
& even.

where  $\tilde{R}_{ss}$  is the discrete Fourier transform of  $R_{ss}$ :

$$\tilde{x}[p] = \mathcal{F}\{x[k]\} = \Delta t \sum_{k=0}^{N-1} x[k] e^{-2\pi i pk/N}$$

$$x[k] = \mathcal{F}^{-1}\{\tilde{x}[p]\} = \Delta f \sum_{p=0}^{N-1} \tilde{x}[p] e^{2\pi i pk/N}$$

$\Delta f = \frac{1}{N \Delta t}$

$= \frac{1}{T}$

Wiener-Khinchin theorem: the F.T. of the autocorrelation function of a WSS stochastic process is equal to the power spectral density of the process  $S_n$ :

$$S_n[p] = \Delta t \sum_{k=0}^{N-1} R_{ss}[k] e^{-2\pi i kp/N}$$

$$\Rightarrow \lambda_p = \frac{S_n[p]}{2\Delta t}$$

$\Leftrightarrow$

Construct matrix of eigenvectors  $U[k, p] = \frac{1}{\sqrt{N}} e^{-2\pi i kp/N}$

$\rightarrow U$  is unitary ( $U^*U = I$ )

& eigenval.  $\Delta E[l, p] = \lambda_p S_{lp}$ .

to get:

$$\sum_j [j, k] = \frac{2\Delta t}{N} \sum_{p=0}^{N-1} \frac{e^{-2\pi i jp/N} e^{2\pi i kp/N}}{S_n[p]}$$

b/c  
 $S_n$  symm  
around  
 $N/2$

$$= 2\Delta f (\Delta t)^2 \sum_{p=0}^{N/2-1} (e^{-2\pi i (j-k)p/N} + e^{2\pi i (j-k)p/N}) S_n[p]$$

Note that b/c  $S$  is real,

$$(\Delta t)^2 \sum_{j, k=0}^{N-1} S[j] S[k] (e^{-2\pi i (j-k)p/N} + e^{2\pi i (j-k)p/N}) = 2 |S|^2 [p]$$

We therefore have:

$$\underline{S}^T \underline{\Sigma}^{-1} \underline{S} = 4 \Delta f \sum_{p=0}^{N/2-1} \frac{|\tilde{s}_p|^2 [p]}{S_n[p]}$$

Note that if we define the inner product

$$\langle \underline{a}, \underline{b} \rangle = 4 \operatorname{Re} \left\{ \Delta f \sum_{p=0}^{N/2-1} \frac{\tilde{a}^*[p] \tilde{b}[p]}{S_n[p]} \right\}$$

Then:

$$\underline{S}^T \underline{\Sigma}^{-1} \underline{S} = \langle \underline{S}, \underline{S} \rangle$$

and:

$$p(\underline{s}_{\text{net}} | n) \propto \exp \left[ -\frac{1}{2} \sum_{d=1}^R \langle \underline{s}_d, \underline{s}_d \rangle \right]$$

What about the signal hypothesis?  $p(\underline{s} | \underline{\lambda}, h) = ?$

$$\underline{s}_{\text{net}} = \underline{n}_{\text{net}} + \underline{h}_{\text{net}}$$

$$\Rightarrow \underline{n}_{\text{net}} = \underline{s}_{\text{net}} - \underline{h}_{\text{net}}$$

$$\Rightarrow p(\underline{s}_{\text{net}} | \underline{\lambda}, h) \propto \exp \left[ -\frac{1}{2} \sum_d \langle \underline{s}_d - \underline{h}_d(\underline{\lambda}), \underline{s}_d - \underline{h}_d(\underline{\lambda}) \rangle \right]$$

## Signal-to-noise ratio (SNR)

Can't do full Bayesian inference on all times

- too computationally expensive
- glitches make significance estimation difficult

Instead a bank of templates is used to identify times when a signal exists. Calculate SNR for each template.

Likelihood ratio:

$$\mathcal{L}(\underline{\lambda}) \equiv \frac{p(s|\underline{\lambda}, h)}{p(s|n)}$$

$$= \frac{\exp\left[\sum_d \langle h_d(\underline{\lambda}), s_d \rangle - \frac{1}{2} \langle h_d(\underline{\lambda}), h_d(\underline{\lambda}) \rangle - \frac{1}{2} \langle s_d, s_d \rangle\right]}{\exp\left[-\frac{1}{2} \sum_d \langle s_d, s_d \rangle\right]}$$

$$\Rightarrow \log \mathcal{L}(\underline{\lambda}) = \sum_d \langle h_d(\underline{\lambda}), s_d \rangle - \frac{1}{2} \langle h_d(\underline{\lambda}), h_d(\underline{\lambda}) \rangle$$

Signals will have some unknown amplitude A & phase  $\phi$ .  
 Don't really care about these if we just want to know when a signal exists.

Amplitude: GW amplitude is inversely proportional to luminosity distance  $D$ . Generate a template  $\underline{h}$  at some fiducial distance  $D_0$  (usually chosen to be 1 Mpc). Then for a signal  $\underline{h}'$  at any distance,

$$\underline{h}' = \frac{D_0}{D} \underline{h} = A \underline{h} \quad \text{and:}$$

$$\log Z = A \langle \underline{h}, \underline{s} \rangle - \frac{1}{2} A^2 \langle \underline{h}, \underline{h} \rangle$$

Maximizing over  $A$ :

Notes: we only consider a single detector here.

$$\frac{d \log Z}{d A} = 0 \Rightarrow A_{\max} = \frac{\langle \underline{h}, \underline{s} \rangle}{\langle \underline{h}, \underline{h} \rangle}$$

$$\Rightarrow \max_A \log Z = \frac{1}{2} \frac{\langle \underline{h}, \underline{s} \rangle^2}{\langle \underline{h}, \underline{h} \rangle} = \frac{f^2}{2}, \quad \text{where}$$

$$f = \frac{|\langle \underline{h}, \underline{s} \rangle|}{\sqrt{\langle \underline{h}, \underline{h} \rangle}}$$

is the SNR.

Phase: Let  $\underline{h} = \underline{h}_0 \cos \phi + \underline{h}_{\pi/2} \sin \phi$ , where

$$\underline{h}(0, \pi/2) [t] = (R_e, I_m) \int_{-\infty}^{\infty} \hat{h}(f) e^{2\pi i f t} df$$

is the waveform with  $\phi=0$  &  $\phi=\pi/2$ , respectively. Then

$$\max_{\phi} \log Z = \frac{\langle \underline{h}_0, \underline{s} \rangle \cos \phi + \langle \underline{h}_{\pi/2}, \underline{s} \rangle \sin \phi}{\sqrt{2 \langle \underline{h}_0, \underline{h}_0 \rangle}}$$

Maximizing over  $\phi$  gives:

$$\phi_{\max} = \arctan\left(\frac{\langle \underline{h}_{\pi/2}, \underline{s} \rangle}{\langle \underline{h}_0, \underline{s} \rangle}\right)$$

Substituting back yields

$$\max_{\phi} g = \left[ \frac{\langle \underline{h}_0, \underline{s} \rangle^2 + \langle \underline{h}_{\pi/2}, \underline{s} \rangle^2}{\langle \underline{h}_0, \underline{h}_0 \rangle} \right]^{\frac{1}{2}}$$

Note that if redefine the inner product as:

$$\langle \underline{h}, \underline{s} \rangle = 4 \int \frac{\underline{h}^*(f) \underline{s}(f)}{S_n(f)} df \quad (\text{i.e., let it be complex})$$

Then:

$$\max_{\phi} g = \frac{|\langle \underline{h}, \underline{s} \rangle|}{\sqrt{\langle \underline{h}, \underline{h} \rangle}} \quad \& \quad \phi_{\max} = \tan^{-1}\left(\frac{\text{Im} \langle \underline{h}, \underline{s} \rangle}{\text{Re} \langle \underline{h}, \underline{s} \rangle}\right)$$

=

Optimal SNR:

Since noise has zero mean, on average the SNR (assuming the template & signal are the same) will be:

$$\bar{g} = \frac{|\langle \underline{h}, \underline{h} + \underline{n} \rangle|}{\sqrt{\langle \underline{h}, \underline{h} \rangle}} = \frac{|\langle \underline{h}, \underline{h} \rangle|}{\sqrt{\langle \underline{h}, \underline{h} \rangle}}$$

$$\Rightarrow \boxed{g_{\text{opt}} = \sqrt{\langle \underline{h}, \underline{h} \rangle}} \text{ is the "optimal" SNR}$$

Overlap: Say the template & signal are not the same - there's some mismatch. Then, on average the SNR will be:

$$\bar{\rho} = \frac{\langle \underline{h}, \underline{h}' \rangle}{\sqrt{\langle \underline{h}, \underline{h} \rangle}}$$

Relative to the optimal SNR for  $\underline{h}'$ , this is:

$$\frac{s}{s_{opt}} = \frac{\langle \underline{h}, \underline{h}' \rangle}{\sqrt{\langle \underline{h}, \underline{h} \rangle \langle \underline{h}', \underline{h}' \rangle}} \equiv \text{"overlap"}$$

$$\begin{aligned} \hat{h} &= \frac{h}{\sqrt{2h_h h_h}} \\ \hat{h}' &= \frac{h'}{\sqrt{2h'_h h'_h}} \\ s/s_{opt} & \end{aligned}$$

Whitening: Can write the inner product as:

$$\langle a, b \rangle = \int_{-\infty}^{\infty} \frac{\hat{a}^*(f)}{\sqrt{S_n(f)}} \frac{\hat{b}(f)}{\sqrt{S_n(f)}} df$$

$$a_w(t) = \int \frac{\hat{a}^*(f)}{\sqrt{S_n(f)}} e^{j2\pi f t} df \quad \text{is the "whitened" time series}$$

$$\langle a, b \rangle = \tilde{\tau}^{-1} \{ \hat{a}_w \cdot \hat{b}_w \}_{(0)} = a_w * b_w = \int a(\tau) b_w(\tau - \omega) d\tau$$

convolution theorem

Time: Can do a trick to get SNR at all times  $T$ .

$$\begin{aligned}\mathcal{F}\{h(t+\tau)\} &= \int h(t+\tau) e^{-2\pi i f t} dt & t' = t + \tau \\ &= \left( \int h(t') e^{-2\pi i f t'} dt' \right) e^{-2\pi i f \tau} & dt' = dt \\ &= \hat{h}(f) e^{-2\pi i f \tau}\end{aligned}$$

Then, the SNR at some time offset  $\tau$  is:

$$\begin{aligned}g(\tau) &= \frac{|\langle h(t-\tau), s \rangle|}{\sqrt{\langle h, h \rangle}} = \frac{1}{\sqrt{\langle h, h \rangle}} \left| \int_{-\infty}^{\infty} \frac{\hat{h}^*(f) \hat{s}(f)}{S_n(f)} e^{2\pi i f \tau} df \right| \\ &= \frac{1}{\sqrt{\langle h, h \rangle}} |\mathcal{F}^{-1}\{\hat{z}\}(\tau)|, \quad \hat{z}(f) = \frac{\hat{h}^*(f) \hat{s}(f)}{S_n(f)}\end{aligned}$$

i.e. can just take the inverse Fourier transform of  $\hat{z}$  to get SNR at all times.

$\Rightarrow N \log N$  operations instead of  $N^2$